## Optimal design for models with correlated observations

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The statement of problem	The multiplicative algorithm
Consider the model $y_j = y_j(t_j) = \theta_1 f_1(t_j) + \ldots + \theta_m f_m(t_j) + \varepsilon_j$ where $t_j \in [-T, T], j = 1,, N$ and $\mathbf{E}\varepsilon_j \varepsilon_i = \sigma^2 \rho(t_j - t_i)$ . For the estimate $\hat{\theta}_{OLS} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T Y$	In general, the optimal design can be a discrete measure, or a continuous density or a combination of these two types. The multiplicative algorithm constructs a discrete design which is very close to the optimal design. Let $\xi^{(r)} = \{x_1, \ldots, x_n; w_1^{(r)}, \ldots, w_n^{(r)}\}$ be a design at the iteration $r$ . Assume that $x_1, \ldots, x_n$ is a rather uniform dense set in the interval $[-1, 1]$ and $w_1^{(0)}, \ldots, w_n^{(0)}$ are nonzero weights, for example, uniform. Following (DPZ2008), we define the updating rule for weights by
the exact design problem has the form $\operatorname{Var}\left(\hat{\theta}_{\mathrm{OLS}}\right) = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \Sigma X (\mathbf{X}^T \mathbf{X})^{-1} \stackrel{\Phi \operatorname{crit}}{\longrightarrow} \min_{t_1, \dots, t_N}.$	$w_i^{(r+1)} = \frac{w_i^{(r)}(\psi(x_i,\xi^{(r)}) - \beta)}{\sum_{j=1}^n w_j^{(r)}(\psi(x_j,\xi^{(r)}) - \beta)},$
Let the design points $\{t_1, \ldots, t_N\}$ be generated by the quantiles of a distribution function,	$i = 1, \ldots, n$ , where $\beta$ is the tuning parameter and

 $v_{iN} = u\left( \begin{pmatrix} i & 1 \end{pmatrix} / \begin{pmatrix} 1 & 1 \end{pmatrix} \right), \ i = 1, \dots, 1, n$ 

where the function  $a : [0, 1] \rightarrow [-T, T]$  is the inverse of a distribution function. Let  $\xi$  be a design measure corresponding to  $a(\cdot)$ .

Under asymptotic settings, the design problem has the form

 $D(\xi) = M^{-1}(\xi)B(\xi,\xi)M^{-1}(\xi) \xrightarrow{\Phi \operatorname{crit}} \min_{\xi}$ 

where  $M(\xi) = \int f(u)f^T(u)\xi(du)$  and  $B(\xi,\nu) = \int \int \rho(u-v)f(u)f^T(v)\xi(du)\nu(dv)$ .

## The optimality condition

Define  $C = \frac{\partial \Phi(D)}{\partial D} = \left(\frac{\partial \Phi(D)}{\partial D_{ij}}\right)_{i,j=1,...,m}$ ,

 $\varphi(x,\xi) = f^T(x) M^{-1}(\xi) B(\xi,\xi) M^{-1}(\xi) C(\xi) M^{-1}(\xi) f(x) ,$ 

 $b(x,\xi) = f^T(x)B^{-1}(\xi,\xi) \int \rho(u-x)f(u)\xi(u),$ 

Theorem 1. Let  $\xi^*$  be any design minimizing the functional  $\Phi(D(\xi)).$  Then the inequality  $\varphi(x,\xi^*) \le b(x,\xi^*)$ 

holds for all  $x \in \mathcal{X}$ . Moreover, there is equality for  $\xi^*$ -almost all x.

 $\psi(x,\xi) = \frac{d(x,\xi)}{b(x,\xi)}.$ 

Note that the optimality condition takes the form  $\psi(x, \xi^*) \leq 1$ .

## Efficiencies of the uniform and arcsine designs

Let us study the *D*-efficiency of the uniform design and the arcsine design for the linear regression model with  $f(x) = (1, x, \dots, x^{m-1})^T$  and different correlation functions.

We determine the *D*-efficiency as

$$\operatorname{Eff}(\xi) = \left(\frac{\det D(\xi^*)}{\det D(\xi)}\right)^{1/m}$$

In our numerical study we computed the *D*-optimal design  $\xi^*$  by the multiplicative algorithm. In Tables 1–3 we can observe that the efficiency of the arcsine design is mainly larger than the efficiency of the uniform design. Moreover, the difference between efficiencies of the arcsine design and the uniform design increases as *m* increases. In addition, the efficiency of the uniform design and the arcsine design decreases as *m* increases.

Table 1. *D*-Efficiencies of the uniform design  $\xi_u$  and the arcsine design  $\xi_a$  for the model with  $f(x) = (1, x, \dots, x^{m-1})^T$  and the exponential correlation function  $\rho(x) = e^{-\lambda |x|}$ .

		$\lambda$	0.5	1.5	2.5	3.5	4.5	5.5
γ	n = 1	$\operatorname{Eff}(\xi_u)$	0.913	0.888	0.903	0.919	0.933	0.944
		$\mathrm{Eff}(\xi_a)$	0.966	0.979	0.987	0.980	0.968	0.954
$\gamma$	n=2	$\operatorname{Eff}(\xi_u)$	0.857	0.832	0.847	0.867	0.886	0.901
		$\mathrm{Eff}(\xi_a)$	0.942	0.954	0.970	0.975	0.973	0.966
$\gamma$	n = 3	$\operatorname{Eff}(\xi_u)$	0.832	0.816	0.826	0.842	0.860	0.876
		$\mathrm{Eff}(\xi_a)$	0.934	0.938	0.954	0.968	0.976	0.981
$\gamma$	n = 4	$\operatorname{Eff}(\xi_u)$	0.826	0.818	0.823	0.835	0.849	0.864
		$\mathrm{Eff}(\xi_a)$	0.934	0.936	0.945	0.957	0.967	0.975

**Theorem 2.** Let  $\xi^*$  be any *D*-optimal design. Then for all  $x \in \mathcal{X}$  we have

 $d(x,\xi^*) \le b(x,\xi^*)$ 

where the functions d is defined by

 $d(x,\xi) = f^{T}(x)M^{-1}(\xi)f(x).$ 

Moreover, there is equality for  $\xi^*$ -almost all x.



Figure 1. The functions  $b(x,\xi)$  and  $d(x,\xi)$  for the regression model with  $f(x) = (1, x, x^2)^T$ and the covariance kernels  $\rho(u-v) = e^{-|u-v|}$  (left),  $\rho(u-v) = \max(0, 1-|u-v|)$  (middle) and  $\rho(u-v) = -\log(u-v)^2$  (right), and the arcsine design  $\xi$ . Table 2. *D*-Efficiencies of the uniform design  $\xi_u$  and the arcsine design  $\xi_a$  for the model with  $f(x) = (1, x, \dots, x^{m-1})^T$  and the triangular corr. function  $\rho(x) = \max\{0, 1 - \lambda |x|\}$ .

	$\lambda$	0.5	1.5	2.5	3.5	4.5	5.5
$\boxed{m=1}$	Eff $(\xi_u)$	0.761	0.866	0.916	0.942	0.956	0.966
	$\operatorname{Eff}(\xi_a)$	0.852	0.941	0.922	0.898	0.874	0.854
m = 2	Eff $(\xi_u)$	0.805	0.777	0.846	0.890	0.916	0.935
	$\operatorname{Eff}(\xi_a)$	0.894	0.907	0.916	0.907	0.890	0.874
m = 3	Eff $(\xi_u)$	0.835	0.756	0.808	0.853	0.884	0.908
	$\operatorname{Eff}(\xi_a)$	0.934	0.908	0.929	0.938	0.936	0.932
$\boxed{m=4}$	Eff $(\xi_u)$	0.821	0.728	0.790	0.837	0.866	0.890
	$\operatorname{Eff}(\xi_a)$	0.931	0.868	0.924	0.947	0.950	0.951

Table 3. *D*-Efficiencies of the uniform design  $\xi_u$  and the arcsine design  $\xi_a$  for the model with  $f(x) = (1, x, \dots, x^{m-1})^T$  and the gaussian correlation function  $\rho(x) = e^{-\lambda x^2}$ .

## Universally optimal designs

**Theorem 3.** Consider the polynomial regression model with  $f(x) = (1, x, x^2, ..., x^{m-1})^T$ ,  $x \in [-1, 1]$ , and the covariance function is  $\rho(x) = \gamma - \beta \ln x^2$  with  $\gamma \ge 0, \beta > 0$ . Then the design with the arcsine density satisfies the necessary conditions for universal optimality.

**Theorem 4.** Consider the polynomial regression model with  $f(x) = (1, x, x^2, ..., x^{m-1})^T$ ,  $x \in [-1, 1]$ , and the covariance function is  $\rho(x) = \gamma + \beta/|x|^{\alpha}$  with  $\gamma \ge 0, \beta > 0$ . Then the design with generalized arcsine density  $p_{\alpha}(x) = \frac{(\Gamma(\alpha + \frac{1}{2}))^2}{2^{\alpha}\Gamma(2^{\alpha+1})} (1 - x^2)^{\alpha - 1/2}, x \in [-1, 1]$ . satisfies the necessary conditions for universal optimality.

Proof. It can be verified that the (generalized) arcsine design satisfies the optimality condition for the *c*-criterion for all *c*.

	$\lambda$	0.5	1.5	2.5	3.5	4.5	5.5
m = 1	Eff $(\xi_u)$	0.758	0.789	0.811	0.830	0.842	0.853
	$\operatorname{Eff}(\xi_a)$	0.841	0.907	0.924	0.932	0.934	0.935
m=2	Eff $(\xi_u)$	0.756	0.698	0.709	0.725	0.739	0.753
	$\operatorname{Eff}(\xi_a)$	0.843	0.833	0.853	0.868	0.877	0.885
m = 3	Eff $(\xi_u)$	0.803	0.662	0.684	0.699	0.711	0.720
	$\operatorname{Eff}(\xi_a)$	0.866	0.771	0.818	0.844	0.859	0.869
m = 4	Eff $(\xi_u)$	0.797	0.630	0.617	0.627	0.648	0.665
	$\operatorname{Eff}(\xi_a)$	0.842	0.713	0.722	0.746	0.776	0.799